

Solutions to Exam 2 (Part II)

1. The random pair (X, Y) has the joint distribution shown in Table 1.

$X \setminus Y$	2	3	4
1	0.1	0	0.1
2	0.3	0.1	0.2
3	0	0.2	0

Table 1: Joint Probability Distribution for X and Y , $p_{(X,Y)}$

- (a) Compute $\Pr(X < Y)$.

Solution:

$$\begin{aligned}
 \Pr(X < Y) &= p_{(X,Y)}(1, 2) + p_{(X,Y)}(1, 3) + p_{(X,Y)}(1, 4) \\
 &\quad p_{(X,Y)}(2, 3) + p_{(X,Y)}(2, 4) + p_{(X,Y)}(3, 4) \\
 &= 0.1 + 0 + 0.1 + 0.1 + 0.2 + 0 \\
 &= 0.5.
 \end{aligned}$$

□

- (b) Compute the marginal distributions of X and Y .

Solution: The marginal distribution of X is

$$p_X(k) = \begin{cases} 0.2, & \text{if } k = 1; \\ 0.6, & \text{if } k = 2; \\ 0.2, & \text{if } k = 3; \\ 0, & \text{elsewhere;} \end{cases}$$

and the marginal distribution of Y is

$$p_Y(k) = \begin{cases} 0.4, & \text{if } k = 2; \\ 0.3, & \text{if } k = 3; \\ 0.3, & \text{if } k = 4; \\ 0, & \text{elsewhere;} \end{cases}$$

□

(c) Show that X and Y are not independent. Give a reason for your answer.

Solution: Note that $p_{(X,Y)}(1, 3) = 0$, while $p_X(1) \cdot p_Y(3) = (0.2) \cdot (0.3)$; so that, $p_X(1) \cdot p_Y(3) = 0.06$. Hence,

$$p_{(X,Y)}(1, 3) \neq p_X(1) \cdot p_Y(3),$$

and, therefore, X and Y are not independent. \square

(d) Compute the expectations $E(X)$, $E(Y)$ and $E(XY)$.

Solution: Compute

$$E(X) = 1 \cdot (0.2) + 2 \cdot (0.6) + 3 \cdot (0.2) = 2.0;$$

$$E(Y) = 2 \cdot (0.4) + 3 \cdot (0.3) + 4 \cdot (0.3) = 2.9;$$

and

$$\begin{aligned} E(XY) &= 1 \cdot 2(0.1) + 1 \cdot 3(0) + 1 \cdot 4(0.1) \\ &\quad + 2 \cdot 2(0.3) + 2 \cdot 3(0.1) + 2 \cdot 4(0.2) \\ &\quad + 3 \cdot 2(0) + 3 \cdot 3(0.2) + 3 \cdot 4(0) \end{aligned}$$

$$\begin{aligned} &= 2(0.1) + 0 + 4(0.1) \\ &\quad + 4(0.3) + 6(0.1) + 8(0.2) \\ &\quad + 0 + 9(0.2) + 0 \end{aligned}$$

$$= 0.2 + 0.4 + 1.2 + 0.6 + 1.8 + 1.6$$

$$= 5.8.$$

\square

(e) Compute the covariance of X and Y .

Solution: Compute

$$\text{Cov}(X, Y) = E(XY) - E(X) \cdot E(Y) = 5.8 - (2.0)(2.9) = 0.$$

\square

2. A random point (X, Y) is distributed uniformly on the triangle with vertices $(0, 0)$, $(4, 0)$ and $(0, 1)$ in the xy -plane.

(a) Give a formula for computing the joint pdf, $f_{(X,Y)}$, of the random vector (X, Y) .

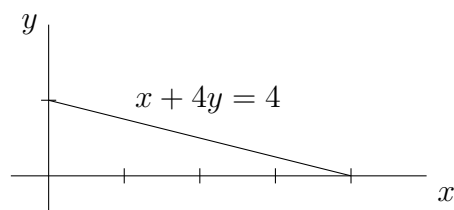


Figure 1: Triangular Region for Problem 2

Solution: Figure 1 shows a sketch of the triangular region with vertices $(0, 0)$, $(4, 0)$ and $(0, 1)$ in the xy -plane. Since the area of the triangle is

$$A = \frac{1}{2}(4)(1) = 2,$$

it follows that the joint pdf of (X, Y) is

$$f_{(X,Y)}(x, y) = \begin{cases} \frac{1}{2}, & \text{if } 0 < x < 4 \text{ and } 0 < y < 1 - x/4; \\ 0, & \text{elsewhere.} \end{cases}$$

□

(b) Compute the marginal distributions f_X and f_Y .

Solution: Compute, for $0 < x < 4$,

$$\begin{aligned} f_X(x) &= \int_{-\infty}^{\infty} f_{(X,Y)}(x, y) dy \\ &= \int_0^{1-x/4} \frac{1}{2} dy \\ &= \frac{1}{2} \left(1 - \frac{x}{4}\right). \end{aligned}$$

Hence,

$$f_X(x) = \begin{cases} \frac{1}{8}(4 - x), & \text{if } 0 < x < 4; \\ 0, & \text{elsewhere.} \end{cases}$$

Similarly, for $0 < y < 1$, compute

$$\begin{aligned} f_Y(y) &= \int_{-\infty}^{\infty} f_{(X,Y)}(x,y) dx \\ &= \int_0^{4-4y} \frac{1}{2} dy \\ &= \frac{1}{2}(4 - 4y); \end{aligned}$$

so that

$$f_Y(y) = \begin{cases} 2(1 - y), & \text{if } 0 < y < 1; \\ 0, & \text{elsewhere.} \end{cases}$$

□

- (c) Are X and Y independent random variables? Give a reason for your answer.

Solution: Observe that

$$f_X(x) \cdot f_Y(y) = \begin{cases} \frac{1}{4}(4 - x)(1 - y), & \text{if } 0 < x < 4 \text{ and } 0 < y < 1; \\ 0, & \text{elsewhere.} \end{cases}$$

Thus, $f_{(X,Y)}(x,y) \neq f_X(x) \cdot f_Y(y)$ for all $(x,y) \in \mathbb{R}^2$. Hence, X and Y are not independent. □

- (d) Compute the expectations $E(X)$, $E(Y)$ and $E(XY)$

Solution: Compute

$$\begin{aligned} E(X) &= \int_{-\infty}^{\infty} x f_X(x) dx \\ &= \int_0^4 \frac{x}{8}(4 - x) dx \\ &= \frac{1}{8} \int_0^4 (4x - x^2) dx \\ &= \frac{1}{8} \left[2x^2 - \frac{x^3}{3} \right]_0^4; \end{aligned}$$

so that

$$E(X) = \frac{4}{3}.$$

Similarly,

$$\begin{aligned} E(Y) &= \int_{-\infty}^{\infty} x f_Y(y) dy \\ &= \int_0^1 2y(1-y) dy \\ &= \int_0^1 (2y - 2y^2) dy \\ &= \left[y^2 - \frac{2}{3}y^3 \right]_0^1; \end{aligned}$$

so that

$$E(Y) = \frac{1}{3}.$$

Next, compute

$$\begin{aligned} E(XY) &= \iint_{\mathbb{R}^2} xy f_{(X,Y)}(x,y) dx dy \\ &= \frac{1}{2} \int_0^1 \int_0^{4(1-y)} xy dx dy \\ &= \frac{1}{4} \int_0^1 [x^2 y]_0^{4(1-y)} dy \\ &= 4 \int_0^1 (1-y)^2 y dy \\ &= 4 \int_0^1 (y - 2y^2 + y^3) dy \\ &= 4 \left[\frac{y^2}{2} - \frac{2}{3}y^3 + \frac{y^4}{4} \right]_0^1; \end{aligned}$$

so that

$$E(XY) = \frac{1}{3}.$$

□

(e) Compute the covariance of X and Y .

Solution: Compute

$$\begin{aligned}\text{Cov}(X, Y) &= E(XY) - E(X) \cdot E(Y) \\ &= \frac{1}{3} - \frac{4}{3} \cdot \frac{1}{3};\end{aligned}$$

so that,

$$\text{Cov}(X, Y) = -\frac{1}{9}.$$

□

3. Suppose that the joint pdf of the random vector (X, Y) is given by

$$f_{(X,Y)}(x, y) = \frac{1}{2\pi} e^{-(x^2+y^2)/2}, \quad \text{for } -\infty < x < \infty \text{ and } -\infty < y < \infty.$$

(a) Verify that $f_{(X,Y)}$ is indeed a joint pdf.

Solution: Use polar coordinates to evaluate

$$\begin{aligned}\iint_{\mathbb{R}^2} f_{(X,Y)}(x, y) \, dx dy &= \frac{1}{2\pi} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-(x^2+y^2)/2} \, dx dy \\ &= \frac{1}{2\pi} \int_0^{2\pi} \int_0^{\infty} e^{-r^2/2} r \, dr d\theta \\ &= \int_0^{\infty} e^{-r^2/2} r \, dr;\end{aligned}$$

thus, making the change of variables $u = r^2/2$,

$$\iint_{\mathbb{R}^2} f_{(X,Y)}(x, y) \, dx dy = \int_0^{\infty} e^{-u} \, du = 1.$$

Hence, $f_{(X,Y)}$ is indeed a joint pdf. □

Alternate Solution: Observe that

$$f_{(X,Y)}(x, y) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2} \cdot \frac{1}{\sqrt{2\pi}} e^{-y^2/2}, \quad (1)$$

for $-\infty < x < \infty$ and $-\infty < y < \infty$, which can be written as

$$f_{(X,Y)}(x,y) = f_X(x) \cdot f_Y(y), \quad \text{for } -\infty < x < \infty \text{ and } -\infty < y < \infty, \quad (2)$$

where f_X and f_Y are both pdfs of a standard normal random variable. It then follows that

$$\iint_{\mathbb{R}^2} f_{(X,Y)}(x,y) \, dx dy = \int_{-\infty}^{\infty} f_X(x) \, dx \int_{-\infty}^{\infty} f_Y(y) \, dy = 1.$$

□

- (b) Compute the marginal distributions f_X and f_Y .

Solution: It follows from (1) that

$$f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}, \quad \text{for } -\infty < x < \infty, \quad (3)$$

and

$$f_Y(y) = \frac{1}{\sqrt{2\pi}} e^{-y^2/2}, \quad \text{for } -\infty < y < \infty. \quad (4)$$

□

- (c) Are X and Y independent? Give a reason for your answer.

Solution: Yes; this follows from (2). □

- (d) Determine the distribution of $X + Y$.

Solution: It follows from (3) and (4) that $X \sim \text{Normal}(0, 1)$ and $Y \sim \text{Normal}(0, 1)$; thus, since X and Y are independent, as seen in the previous part,

$$X + Y \sim \text{Normal}(0, 2). \quad (5)$$

□

- (e) Compute $\Pr(-\sqrt{2} < X + Y \leq 2\sqrt{2})$.

Solution: Set $Z = \frac{X + Y}{\sqrt{2}}$. It then follows from (5) that $Z \sim \text{Normal}(0, 1)$.

Then,

$$\begin{aligned} \Pr(-\sqrt{2} < X + Y \leq 2\sqrt{2}) &= \Pr\left(-1 < \frac{X + Y}{\sqrt{2}} \leq 2\right) \\ &= \Pr(-1 < Z \leq 2), \end{aligned}$$

where $Z \sim \text{Normal}(0, 1)$. Consequently,

$$\begin{aligned}\Pr(-\sqrt{2} < X + Y \leq 2\sqrt{2}) &= F_Z(2) - F_Z(-1) \\ &= F_Z(2) - (1 - F_Z(1)) \\ &= F_Z(2) + F_Z(1) - 1.\end{aligned}$$

We then have that

$$\Pr(-\sqrt{2} < X + Y \leq 2\sqrt{2}) \approx 0.9772 + 0.8413 - 1 = 0.8185.$$

□