Solutions to Review Problems for Exam 1

1. There are 5 red chips and 3 blue chips in a bowl. The red chips are numbered 1, 2, 3, 4, 5 respectively, and the blue chips are numbered 1, 2, 3 respectively. If two chips are to be drawn at random and without replacement, find the probability that these chips are have either the same number or the same color.

Solution: Let R denote the event that the two chips are red. Then the assumption that the chips are drawn at random and without replacement implies that

$$\Pr(R) = \frac{\binom{5}{2}}{\binom{8}{2}} = \frac{5}{14}.$$

Similarly, if B denotes the event that both chips are blue, then

$$\Pr(B) = \frac{\binom{3}{2}}{\binom{8}{2}} = \frac{3}{28}$$

It then follows that the probability that both chips are of the same color is

$$\Pr(R \cup B) = \Pr(R) + \Pr(B) = \frac{13}{28},$$

since R and B are mutually exclusive.

Let N denote the event that both chips show the same number. Then,

$$\Pr(N) = \frac{3}{\binom{8}{2}} = \frac{3}{28}.$$

Finally, since $R \cup B$ and N are mutually exclusive, then the probability that the chips are have either the same number or the same color is

$$\Pr(R \cup B \cup N) = \Pr(R \cup B) + \Pr(N) = \frac{13}{28} + \frac{3}{28} = \frac{16}{28} = \frac{2}{7}.$$

2. A person has purchased 10 of 1,000 tickets sold in a certain raffle. To determine the five prize winners, 5 tickets are drawn at random and without replacement. Compute the probability that this person will win at least one prize.

Solution: Let N denote the event that the person will not win any prize. Then

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$$\Pr(N) = \frac{\binom{995}{10}}{\binom{1000}{10}};$$
(1)

that is, the probability of purchasing 10 non-winning tickets. It follows from (1) that

$$Pr(N) = \frac{(990)(989)(988)(987)(986)}{(1000)(999)(998)(997)(996)}$$
$$= \frac{435841667261}{458349513900}$$
(2)
$$\approx 0.9509.$$

Thus, using the result in (2), the probability of the person winning at least one of the prizes is

$$Pr(N^c) = 1 - Pr(N)$$

 $\approx 1 - 0.9509$
 $= 0.0491,$

or about 4.91%.

3. Let $(\mathcal{C}, \mathcal{B}, \Pr)$ denote a probability space, and let E_1, E_2 and E_3 be mutually exclusive events in \mathcal{B} . Find $\Pr[(E_1 \cup E_2) \cap E_3]$ and $\Pr(E_1^c \cup E_2^c)$.

Solution: Since E_1 , E_2 and E_3 are mutually disjoint events, it follows that $(E_1 \cup E_2) \cap E_3 = \emptyset$; so that

$$\Pr[(E_1 \cup E_2) \cap E_3] = 0.$$

Next, use De Morgan's law to compute

$$Pr(E_1^c \cup E_2^c) = Pr([E_1 \cap E_2]^c)$$
$$= Pr(\emptyset^c)$$
$$= Pr(\mathcal{C})$$
$$= 1.$$

4. Let $(\mathcal{C}, \mathcal{B}, \Pr)$ denote a probability space, and let A and B events in \mathcal{B} . Show that

$$\Pr(A \cap B) \le \Pr(A) \le \Pr(A \cup B) \le \Pr(A) + \Pr(B).$$
(3)

Solution: Since $A \cap B \subseteq A$, it follows that

$$\Pr(A \cap B) \leqslant \Pr(A). \tag{4}$$

Similarly, since $A \subseteq A \cup B$, we get that

$$\Pr(A) \leqslant \Pr(A \cup B). \tag{5}$$

Next, use the identity

$$\Pr(A \cup B) = \Pr(A) + \Pr(B) - \Pr(A \cap B),$$

and fact that that

$$\Pr(A \cap B) \ge 0,$$

to obtain that

$$\Pr(A \cup B) \leqslant \Pr(A) + \Pr(B). \tag{6}$$

Finally, combine (4), (5) and (6) to obtain (3).

5. Let $(\mathcal{C}, \mathcal{B}, \Pr)$ denote a probability space, and let E_1 , E_2 and E_3 be mutually independent events in \mathcal{B} with probabilities $\frac{1}{2}$, $\frac{1}{3}$ and $\frac{1}{4}$, respectively. Compute the exact value of $\Pr(E_1 \cup E_2 \cup E_3)$.

Solution: First, use De Morgan's law to compute

$$\Pr[(E_1 \cup E_2 \cup E_3)^c] = \Pr(E_1^c \cap E_2^c \cap E_3^c)$$
(7)

Then, since E_1 , E_2 and E_3 are mutually independent events, it follows from (7) that

$$\Pr[(E_1 \cup E_2 \cup E_3)^c] = \Pr(E_1^c) \cdot \Pr(E_2^c) \cdot \Pr(E_3^c),$$

so that

$$\Pr[(E_1 \cup E_2 \cup E_3)^c] = (1 - \Pr(E_1))(1 - \Pr(E_2))(1 - \Pr(E_3))$$
$$= \left(1 - \frac{1}{2}\right) \left(1 - \frac{1}{3}\right) \left(1 - \frac{1}{4}\right)$$
$$= \frac{1}{2} \cdot \frac{2}{3} \cdot \frac{3}{4},$$

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so that

$$\Pr[(E_1 \cup E_2 \cup E_3)^c] = \frac{1}{4}.$$
 (8)

It then follows from (8) that

$$\Pr(E_1 \cup E_2 \cup E_3) = 1 - \Pr[(E_1 \cup E_2 \cup E_3)^c] = \frac{3}{4}.$$

6. Let $(\mathcal{C}, \mathcal{B}, \Pr)$ denote a probability space, and let E_1 , E_2 and E_3 be mutually independent events in \mathcal{B} with $\Pr(E_1) = \Pr(E_2) = \Pr(E_3) = \frac{1}{4}$. Compute $\Pr[(E_1^c \cap E_2^c) \cup E_3]$.

Solution: First, use De Morgan's law to compute

$$\Pr[((E_1^c \cap E_2^c) \cup E_3)^c] = \Pr[(E_1^c \cap E_2^c)^c \cap E_3^c]$$
(9)

Next, use the assumption that E_1 , E_2 and E_3 are mutually independent events to obtain from (9) that

$$\Pr[((E_1^c \cap E_2^c) \cup E_3)^c] = \Pr[(E_1^c \cap E_2^c)^c] \cdot \Pr[E_3^c],$$
(10)

where

$$\Pr[E_3^c] = 1 - \Pr(E_3) = \frac{3}{4},\tag{11}$$

and

$$\Pr[(E_1^c \cap E_2^c)^c] = 1 - \Pr[E_1^c \cap E_2^c] = 1 - \Pr[E_1^c] \cdot \Pr[E_2^c],$$
(12)

by the independence of E_1 and E_2 .

It follows from the calculations in (12) that

$$\Pr[(E_1^c \cap E_2^c)^c] = 1 - (1 - \Pr[E_1])(1 - \Pr[E_2])$$

= $1 - \left(1 - \frac{1}{4}\right)\left(1 - \frac{1}{4}\right)$
= $1 - \frac{3}{4} \cdot \frac{3}{4}$
= $\frac{7}{16}$ (13)

Substitute (11) and the result of the calculations in (13) into (10) to obtain

$$\Pr[((E_1^c \cap E_2^c) \cup E_3)^c] = \frac{7}{16} \cdot \frac{3}{4} = \frac{21}{64}.$$
 (14)

Finally, use the result in (14) to compute

$$\Pr[(E_1^c \cap E_2^c) \cup E_3^c] = 1 - \Pr[((E_1^c \cap E_2^c) \cup E_3)^c]$$
$$= 1 - \frac{21}{64}$$
$$= \frac{43}{64}.$$

- 7. A machine produces parts that are either good (90%), slightly defective (2%), or obviously defective (8%). Produced parts get passed through an automatic inspection machine, which is able to detect any part that is obviously defective and discard it.
 - (a) If a part passes the inspection, what is the probability that is is a good part?

Solution: Let G denote the event that the machine produces a good part, S denote the event that the machine produces a slightly defective part, and D the event that the machine produces an obviously defective part. We are then given that

 $\Pr(G) = 0.90, \quad \Pr(S) = 0.02 \quad \text{and} \quad \Pr(D) = 0.08.$

A part passes inspection if it is good part or if it is slightly defective; in other words, if the complement of event D occurs (note that $D^c = G \cup S$). Thus, the probability that a part is good, given that it passed inspection is the conditional probability

$$Pr(G \mid D^{c}) = \frac{Pr(G \cap D^{c})}{Pr(D^{c})}$$
$$= \frac{Pr(G)}{Pr(G \cup S)}$$
$$= \frac{0.90}{0.92}$$
$$= \frac{45}{46}.$$

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(b) Given that a part passes the inspection, what is the probability that it is slightly defective?

Solution: In this case we compute the conditional probability

$$\Pr(S \mid D^c) = \frac{\Pr(S \cap D^c)}{\Pr(D^c)}$$
$$= \frac{\Pr(S)}{\Pr(G \cup S)}$$
$$= \frac{0.02}{0.92}$$
$$= \frac{1}{46}$$

(c) Assume that a one-year warranty is given for the parts that are shipped to customers. Suppose that a good part fails within the first year with probability 0.01, while a slightly defective part fails within the first year with probability 0.10. What is the probability that a customer receives a part that fails within the first year and is therefore entitled to a warranty replacement?

Solution: Let F denote the event that a part that has passed inspection will fail within the first year after shipping. Let G_I denote the event that a good part has passed inspection and been shipped. From part (a) we have that

$$\Pr(G_I) = \Pr(G \mid D^c) = \frac{45}{46}.$$

Similarly, denoting by S_I the event that a slightly defective part has passed inspection, we have from part (b) that

$$\Pr(S_I) = \Pr(S \mid D^c) = \frac{1}{46}.$$

We are given that

$$\Pr(F \mid G_I) = 0.01$$
 and $\Pr(F \mid S_I) = 0.10$.

It then follows from the Law of Total Probability that

$$Pr(F) = Pr(G_I) \cdot Pr(F \mid G_I) + Pr(S_I) \cdot Pr(F \mid S_I)$$
$$= \frac{45}{46} \cdot (0.01) + \frac{1}{46} \cdot (0.10)$$
$$\doteq 0.0112.$$

Thus, the probability that a customer receives a part that fails within the first year is about 1.12%.

8. Toss a fair coin three times in a row. Let A denote the event that either the three tosses yield three heads or three tails; B the event that at least two heads come up; and C the event that at most two tails come up. Out of the pairs of events: (A, B), (A, C), and (B, C), determine the ones that are independent and the ones that are dependent. Explain your reasoning.

Solution: The sample space for this experiment is

$$\mathcal{C} = \{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\}.$$

The events A, B and C are

$$A = \{HHH, TTT\},\$$
$$B = \{HHH, HHT, HTH, THH\},\$$

and

$$C = \{HHH, HHT, HTH, HTT, THH, THT, TTH\},\$$

respectively. Since all the elements of \mathcal{C} are equally likely, it follows that

$$\Pr(A) = \frac{1}{4}, \quad \Pr(B) = \frac{1}{2}, \quad \text{and} \quad \Pr(C) = \frac{7}{8}.$$

Note that $A \cap B = \{HHH\}$, so that

$$\Pr(A \cap B) = \frac{1}{8} = \frac{1}{4} \cdot \frac{1}{2} = \Pr(A) \cdot \Pr(B);$$

thus, A and B are independent.

Next, compute $A \cap C = \{HHH\}$, so that

$$\Pr(A \cap C) = \frac{1}{8} \neq \frac{1}{4} \cdot \frac{7}{8} = \Pr(A) \cdot \Pr(C);$$

thus, A and C are not independent.

Finally, compute $B \cap C = \{HHH, HHT, HTH, THH\}$, so that

$$\Pr(B \cap C) = \frac{1}{2} \neq \frac{1}{2} \cdot \frac{7}{8} = \Pr(B) \cdot \Pr(C);$$

thus, B and C are not independent.

9. A bowl contains 10 chips of the same size and shape. One and only one of these chips is red. Draw chips from the bowl at random, one at a time and without replacement, until the red chip is drawn. Let X denote the number of draws needed to get the red chip. Determine the pmf of X and compute $Pr(X \leq 4)$.

Solution: Compute

$$Pr(X = 1) = \frac{1}{10}$$

$$Pr(X = 2) = \frac{9}{10} \cdot \frac{1}{9} = \frac{1}{10}$$

$$Pr(X = 3) = \frac{9}{10} \cdot \frac{8}{9} \cdot \frac{1}{8} = \frac{1}{10}$$

$$\vdots$$

$$Pr(X = 10) = \frac{1}{10}$$

Thus,

$$p_{X}(k) = \begin{cases} \frac{1}{10} & \text{for } k = 1, 2, \dots, 10; \\ 0 & \text{elsewhere.} \end{cases}$$
(15)

Next, use (15) to compute

$$\Pr(X \le 4) = \sum_{k=1}^{4} p_X(k) = \frac{4}{10} = \frac{2}{5}.$$

10. Let X have pmf given by $p_X(x) = \frac{1}{3}$ for x = 1, 2, 3 and p(x) = 0 elsewhere. Give the pmf of Y = 2X + 1.

Solution: Note that the possible values for Y are 3, 5 and 7 Compute

$$\Pr(Y=3) = \Pr(2X+1=3) = \Pr(X=1) = \frac{1}{3}.$$

Similarly, we get that

$$\Pr(Y=5) = \Pr(X=2) = \frac{1}{3},$$

and

$$\Pr(Y = 7) = \Pr(X = 3) = \frac{1}{3}.$$

Thus,

$$p_{Y}(k) = \begin{cases} \frac{1}{3} & \text{ for } k = 3, 5, 7; \\ \\ 0 & \text{ elsewhere.} \end{cases}$$

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11. Let X have pmf given by $p_X(x) = \left(\frac{1}{2}\right)^x$ for $x = 1, 2, 3, \ldots$ and $p_X(x) = 0$ elsewhere. Give the pmf of $Y = X^3$.

Solution: Compute, for $y = k^3$, for $k = 1, 2, 3, \ldots$,

$$\Pr(Y = y) = \Pr(X^3 = k^3) = \Pr(X = k) = \left(\frac{1}{2}\right)^k,$$

so that

$$\Pr(Y = y) = \left(\frac{1}{2}\right)^{y^{1/3}}, \text{ for } y = k^3, \text{ for some } k = 1, 2, 3, \dots$$

Thus,

$$p_{\scriptscriptstyle Y}(y) = \begin{cases} \left(\frac{1}{2}\right)^{y^{1/3}}, & \text{for } y = k^3, \text{ for some } k = 1, 2, 3, \ldots; \\ \\ 0 & \text{elsewhere.} \end{cases}$$

12. Let $f(x) = \begin{cases} \frac{1}{x^2} & \text{if } 1 < x < \infty; \\ 0 & \text{if } x \le 1, \end{cases}$ and define a probability on the Borel σ -field

of the real line \mathbb{R} by $\Pr[(a, b)] = \int_a^b f(x) \ dx$, for all intervals, (a, b).

If E_1 denote the interval (1,2) and E_2 the interval (4,5), compute $\Pr(E_1)$, $\Pr(E_2)$, $\Pr(E_1 \cup E_2)$ and $\Pr(E_1 \cap E_2)$.

Solution: Compute

$$\Pr(E_1) = \int_1^2 \frac{1}{x^2} dx = -\frac{1}{x} \Big|_1^2 = \frac{1}{2},$$

$$\Pr(E_2) = \int_4^5 \frac{1}{x^2} dx = -\frac{1}{x} \Big|_4^5 = \frac{1}{20},$$

$$\Pr(E_1 \cup E_2) = \Pr(E_1) + \Pr(E_2) = \frac{11}{20},$$

since E_1 and E_2 are mutually exclusive, and

$$\Pr(E_1 \cap E_2) = 0,$$

since E_1 and E_2 are mutually exclusive.

13. A mode of a distribution of a random discrete variable X is a value of x that maximizes the pmf of X. If there is only one such value, it is called the mode of the distribution.

Let X have pmf given by $p(x) = \left(\frac{1}{2}\right)^x$ for $x = 1, 2, 3, \ldots$, and p(x) = 0 elsewhere. Compute a mode of the distribution.

Solution: Note that p(x) is decreasing; so, p(x) is maximized when x = 1. Thus, 1 is the mode of the distribution of X.

14. Let $f(x) = \begin{cases} cx(1-x), & \text{if } 0 < x < 1; \\ 0 & \text{elsewhere,} \end{cases}$ where c is a positive constant.

(a) Determine the value of c so that $\Pr[(a,b)] = \int_a^b f(x) \, dx$, for all intervals, (a,b), defines a probability on the Borel σ -field of the real line \mathbb{R} .

Solution: We choose c so that $Pr(\mathbb{R}) = 1$, where

$$\Pr(\mathbb{R}) = \int_{-\infty}^{\infty} f(x) dx$$
$$= \int_{0}^{1} cx(1-x) dx$$
$$= c \int_{0}^{1} [x-x^{2}] dx$$
$$= c \left[\frac{x^{2}}{2} - \frac{x^{3}}{3}\right]_{0}^{1}$$
$$= \frac{c}{6}.$$

It then follows that c = 6.

(b) For each $x \in \mathbb{R}$, define $F(x) = \Pr[(-\infty, x]]$. Compute F and sketch its graph. Find the value of x for which F(x) = 0.5.

Solution: We compute $F(x) = \int_{-\infty}^{x} f(t) dt$, for $x \in \mathbb{R}$, where

$$f(t) = \begin{cases} 6t(1-t), & \text{if } 0 < t < 1; \\ 0, & \text{elsewhere.} \end{cases}$$

If $x \leq 0$ we have that f(t) = 0 for all $t \leq x$, so that

$$F(x) = 0, \quad \text{for } x \leq 0.$$

If $0 < x \leq 1$, we get that

$$F(x) = \int_0^x 6t(1-t) dt$$

= $6\left[\frac{t^2}{2} - \frac{t^3}{3}\right]_0^x$
= $3x^2 - 2x^3$.

Finally, if $x \ge 1$, we have that

$$F(x) = 1.$$

Putting all these calculations together we get

$$F(x) = \begin{cases} 0, & \text{if } x \leq 0; \\ 3x^2 - 2x^3, & \text{if } 0 < x \leq 1; \\ 1, & \text{if } x > 1. \end{cases}$$

A sketch of the graph of F is found in Figure 1.



Figure 1: Sketch of F(x) in Problem 14

Observe that $F(0.5) = 3\left(\frac{1}{2}\right)^2 - 2\left(\frac{1}{2}\right)^3 = \frac{3}{4} - \frac{1}{4} = \frac{1}{2}$, so that $x = \frac{1}{2}$ is the unique value of x for which F(x) = 0.5.